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Efficacy and safety of corticosteroid regimens for the treatment of hospitalized COVID-19 patients: a meta-analysis

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Aim: To evaluate the efficacy and safety of corticosteroids for treating hospitalized COVID-19 patients. **Materials & methods:** Efficacy outcomes included time to negative SARS-CoV-2 tests, length of stay, duration and incidence of intensive unit care stay, incidence of mortality and duration and incidence of mechanical ventilation. Safety outcomes included the incidence of adverse events and severe adverse events, incidence of hyperglycemia and incidence of nosocomial infections. **Results:** Ninety-five randomized controlled trials (RCTs) and observational studies (n = 42,205) were included. Corticosteroids were associated with increased length of stay (based on RCT only), increased time to negative tests, decreased length of mechanical ventilation and increased odds of hyperglycemia. **Conclusion:** Corticosteroids should be considered in patients requiring mechanical ventilation, and glycemic monitoring may be needed when administering corticosteroids.

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Keywords: corticosteroids • COVID-19 • length of stay • mechanical ventilation • mortality • SARS-CoV-2

In March 2020 the WHO designated COVID-19, which is caused by SARS-CoV-2, as a global pandemic [1–4]. The spread of SARS-CoV-2 has been considered to be a severe public health crisis due to its high transmission, hospitalization and mortality rates [5–8]. One of the largest factors attributed to the high mortality rate associated with COVID-19 was the depletion of hospital resources as a result of surges in cases [9–11]. The care of critically ill COVID-19 patients require valuable and scarce supplies such as ICU beds, ventilators, personal protective equipment and hospital staff [12]; thus, physicians have sometimes been forced to triage critically ill COVID-19 patients [13] and divert resources from other specialties such as oncology, leading to the rescheduling of appointments for non-COVID patients [14].

Recent advancements in SARS-CoV-2 vaccination programs, including the US FDA-approved Comirnaty vaccine produced by Pfizer-BioNTech [15], aims to curb the transmission of the virus, induce herd immunity and relieve the pressure of the pandemic on hospital resources. However, new viral variants of interest (VOIs), such as Lambda and Delta [16], threaten to introduce new COVID-19 surges in the coming months with their increased infectivity [17], immune resistance [17,18] and reduced sensitivity to antibody neutralization [19]. These newly discovered variants, in addition to widespread vaccine hesitancy movements [20] and poor vaccine accessibility in low- to middle-income countries (LMICs) [21] are likely to extend the pandemic well into 2022 [22]. Therefore, further investigations are still required to determine the optimal treatment regimen for the management of COVID-19 surges in the coming months.



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Currently, most of the proposed COVID-19 treatment regimens consist of repurposed pharmacological compounds that were initially chosen based on their antiviral efficacy against SARS-CoV-2 in vitro or in small observational studies. However, many repurposed regimens such as hydroxychloroquine and lopinavir/ritonavir have been consistently shown in randomized controlled trials (RCTs) and meta-analyses to confer no efficacy against SARS-CoV-2 in clinical settings [23-26]. In early 2020, a series of observational and prognostic studies brought a new category of repurposed therapies, corticosteroids, into the limelight [27]. These early findings are corroborated by encouraging results from the RECOVERY trial, which is the world's largest clinical trial for COVID-19 treatments to date [28]. The RECOVERY trial concluded that dexamethasone may confer survival benefits among patients receiving mechanical ventilation. Unfortunately, no other authoritative RCTs, including the REMAP-CAP [29] and CoDEX trial [30], were able to identify a significant beneficial effect associated with corticosteroids. This was partially due to insufficient sample sizes as a result of premature trial terminations following the publication of results from the RECOVERY trial [31]. Additionally, pre-pandemic studies of corticosteroids in other viral pneumonias, including the original 2003 SARS-CoV [32], the Middle East respiratory syndrome (MERS) [33] and influenza [34], have associated the use of corticosteroids with delayed viral clearance and increased mortality. These conflicting findings call into question the efficacy of corticosteroids, prompting further investigations into this category of treatments.

Biologically, the rationale for the use of corticosteroids in the treatment of COVID-19 is to reduce excessive host inflammatory responses in the lungs through the inflammatory phase of the disease, during which immunopathological factors play a bigger role in disease progression compared with viral replication [35,36]. For some patients, impaired interferon responses [37,38] and the activation of intracellular antiviral pathways triggered by viral RNA may lead to symptoms that mimic macrophage activation syndrome (MAS) during the inflammatory phase [39]. MAS is associated with impaired cytolytic capabilities of CD8⁺ T cells and natural killer (NK) cells [40,41], which are normally tasked with lysing infected cells to prevent viral replication and the excessive secretion of inflammatory cytokines [42]. This impairment is driven by high levels of pro-inflammatory cytokine IL-6, which forms a vicious cycle of cytokine-driven cytokine secretion [43] leading to cytokine storms [39]. IL-6 is also implicated in the excessive differentiation of Th2 cells and Th17 cells while inhibiting the differentiation of antiviral Th1 cells [40,44–46]. This action prevents the host immune system from mounting an effective antiviral response while inducing a nonproductive, eosinophilic immune response [47]. COVID-19 patients with elevated IL-6 levels and eosinophilic immune responses have been associated with the development of severe lung injuries and acute respiratory distress syndrome (ARDS) [48,49].

As one of the most commonly used therapies for cytokine storms and MAS in patients with autoimmune and inflammatory diseases, corticosteroids appear to be the ideal regimen for treating COVID-related hyperinflammatory disorders [50]. With appropriate administration times and dosages, it is proposed that corticosteroids can reduce the production of proinflammatory cytokines and inhibit overactive immune cells, thus preventing lung damage and ARDS [51–53]. On the other hand, corticosteroids' immunosuppressive properties may impair viral clearance and increase the rate of dangerous nosocomial infections [54]. It is evident that the role of corticosteroids requires further clarification to examine whether its immunomodulatory efficacy outweighs its adverse effects. Thus, the current meta-analysis assessed the efficacy and safety of corticosteroids compared with standard of care, and studied the impact of factors such as dosage and time of administration on patient-important outcomes.

Methods

We conducted this systematic review and meta-analysis in accordance with the PRISMA 2020 guidelines [55]. The completed PRISMA 2020 checklist is provided in Supplementary Table 1. This review was prospectively registered on PROSPERO (CRD42021233108), the international prospective register of systematic reviews [56].

Study identification

To identify relevant studies for inclusion in this review, we created an English search strategy based on databasespecific COVID-19 search strings provided by the Rudolph Matas Library of the Health Sciences, Tulane University (LA, USA) [57] in combination with treatment-relevant keywords such as 'dexamethasone', 'prednisone', 'methylprednisolone', 'corticosteroid', 'glucocorticoid' and 'steroid' etc. Using this strategy, we searched the following databases from 1 January 2020 to 12 July 2021: Medical Literature Analysis and Retrieval System Online (MED-LINE), Excerpta Medica Database (EMBASE), and PubMed. Additionally, we utilized a Chinese search strategy equivalent to our English strategy, and systematically searched the following Chinese literature databases from 1 January 2020 to 12 July 2021: Wanfang Data, Wanfang Med Online, SinoMed, China National Knowledge Infrastructure (CNKI) and Chongqing VIP Information (CQVIP). The search strategies used for the database searches are appended in Supplementary Tables 2–9. Last, we hand searched the reference sections of six previous systematic reviews [34,58–62] to identify relevant studies that were not identified by our database searches.

Eligibility criteria

We included both randomized and non-randomized comparative studies that met the following inclusion criteria: compared any corticosteroid therapy with standard of care, or compared any corticosteroid therapy with an adjuvant therapy to adjuvant therapy alone, and included hospitalized COVID-19 patients. We only included studies using adjuvant therapies if the adjuvant therapy between the control and experimental arms are the same in order to minimize the effect of the adjuvant therapies on treatment outcomes.

Outcome measures

Our efficacy outcomes included: time to negative conversion of SARS-CoV-2 tests, overall length of hospitalization, length of ICU hospitalization, incidence of progression to ICU hospitalization, incidence of mortality, incidence of progression to mechanical ventilation and length of mechanical ventilation. The safety outcomes included: incidence of all-cause adverse events, incidence of investigator-defined severe adverse events, incidence of hyperglycemia or corticosteroid-induced diabetes mellitus and incidence of nosocomial bacterial or fungal infections.

Study selection

All title/abstract entries identified from our literature searches were entered into an independent and in duplicate title/abstract screening process via Rayyan (https://www.rayyan.ai/) [63]. Abstract entries that were deemed eligible were then entered into a subsequent in duplicate full-text screening process. During the selection process, all disagreements were referred to a senior author for resolution. The PRISMA flowchart [64] depicting our study selection process is shown in Figure 1.

Data extraction

We performed data extraction independently and in duplicate using an extraction sheet developed *a priori*. Categories of extracted information included: patient demographics and baseline characteristics, descriptions of study methodology, treatment descriptions and outcome measures. The full list of extracted items are listed on our PROSPERO registration page. For observational studies reporting data for both matched and unmatched cohorts, we only extracted and analyzed the data from matched cohorts. All disagreements were resolved by recruiting a senior author to review the data.

Missing data

For studies with missing data required for meta-analysis, including mean values for continuous outcomes and measures of variance, we made attempts to contact the corresponding authors to obtain unpublished data. For studies that presented median and interquartile range (IQR) for continuous outcomes, we used methods recommended by Luo *et al.* [65] and Wan *et al.* [66] to estimate the mean and standard deviation (SD) for meta-analysis, given that the data is normally distributed as examined using methods proposed by Shi *et al.* [67]. We did not include skewed median and IQR values in the meta-analysis and elected to describe these results narratively.

Risk of bias assessment

The risk of bias in our included RCTs were assessed independently and in duplicate using version 2 of the Cochrane risk-of-bias tool for randomized trials (RoB2) [68]. The risk of bias of non-randomized comparative studies was assessed in duplicate using the risk of bias in non-randomized studies of interventions (ROBINS-I) tool [69]. All disagreements were referred to a senior author for resolution.

Quality of evidence

We assessed the quality of evidence for our outcomes using the Grading of Recommendations, Assessment, Development and Evaluations (GRADE) framework [70,71]. A summary of the meta-analysis outcomes and GRADE ratings were presented in a GRADE summary of findings table [72] generated using GRADEpro (https://gradepro.org/) [73].

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Figure 1. PRISMA flowchart for the identification and selection of studies.

CNKI: China national knowledge infrastructure; CQVIP: Chongqing VIP information; EMBASE: Excerpta medica database; MEDLINE: Medical literature analysis and retrieval system online; PRISMA: Preferred reporting items for systematic reviews and Meta-Analyses.

Statistical analysis

All statistical analyses were conducted using R v4.1.1 (https://www.r-project.org/) [74]. For outcomes with more than two studies presenting eligible data, we performed random-effects meta-analyses using the meta v4.19 library (https: //cran.r-project.org/web/packages/meta/) [75]. For dichotomous/incidence outcomes, we expressed the treatment effect as odds ratios (ORs) and 95% CIs for data pooling. The treatment effects of continuous outcomes were expressed as mean differences (MDs) and 95% CIs. Results from individual studies and data synthesis were displayed on forest plots.

Zero event studies

For studies reporting zero events in one or both of its treatment arms, we applied treatment arm continuity correction (TACC) [76] to complete the meta-analysis.

Heterogeneity assessment

The presence of heterogeneity was examined using the Cochran's Q test [77] with a significance level of p < 0.10 as recommended by the Cochrane Handbook [78]. Heterogeneity was subsequently quantified using I² statistics [77,79]. We interpreted 30% <I² <75% as moderate heterogeneity and I² \geq 75% as serious heterogeneity, as per recommendations from the Cochrane Handbook [78].

Publication bias & small study effects

We created funnel plots for outcomes with 10 or more included studies to identify small study effects within the meta-analysis as an indication for publication bias [78]. Egger's regression test was used to quantitatively evaluate the presence of asymmetry within the funnel plots [80]. We did not draw funnel plots nor conduct Egger's regression test for outcomes with fewer than 10 included studies as Egger's test does not have sufficient power in these circumstances [81].

Meta-regression & subgroup analyses

We performed meta-regression analyses on the proportion of patients with severe disease (defined as per individual investigators' criteria), mean baseline sequential organ failure assessment (SOFA) scores, and follow-up duration for mortality.

Furthermore, we performed the following subgroup analyses: corticosteroid regimens, corticosteroid dose, with methylprednisolone equivalent dose of $\leq 1.5 \text{ mg/kg/day}$ or $\leq 100 \text{ mg/day}$ defined as low dose and methylprednisolone equivalent dose of > 1.5 mg/kg/day or > 100 mg/day defined as high dose, similar to the threshold set in previous studies [82,83], time from hospitalization to corticosteroid administration, with a time to administration ≤ 3 days defined as early/immediate administration, and a time to administration > 3 days defined as late/delayed administration. Additionally, we performed subgroup analyses by study design (randomized studies vs observational studies) and risk of bias (low/some concerns/moderate risk of bias vs high/serious/critical risk of bias) to examine the impact of different study methodologies.

Results

Included studies

We identified and screened 9865 potentially eligible titles and abstracts (following deduplication using Endnote 20 [https://endnote.com/]). A total of 250 full-text articles were subsequently retrieved and screened. After screening, 10 RCTs [28–30,84–90] and 85 observational studies [83,91–174] were included in this systematic review and meta-analysis (Figure 1), with a total of 42,205 hospitalized COVID-19 patients. In the article selection and data extraction process, we contacted the corresponding authors from seven studies [95,147,149,155,175–177] with unclear or missing data and two authors [149,155] responded to our inquiries. Of the remaining studies, two [95,147] were included with a subset of the reported outcomes without response from authors, and three studies [175–177] were excluded due to a lack of relevant outcomes/insufficient data.

Across all included studies, 37 studies (38.9%) [84,86,89,90,93,94,99,101,104,106,107,112,115,119,121,125,126,130,136,138,142, 143,146,147,152–155,161,162,165,167–170,172,174] included treatment arms that assessed methylprednisolone, 10 studies (10.5%) [28,30,88,100,105,114,117,125,141,171] included treatment arms that assessed dexamethasone, 3 studies (3.2%) [29,85,87] administered hydrocortisone, 1 study (1.1%) [92] administered prednisone and 1 study (1.1%) [145] administered prednisolone. The study by Ko *et al.* [125] had two corticosteroid arms, with one arm receiving dexamethasone and the other arm receiving methylprednisolone. The remaining studies did not specify the exact corticosteroids used and/or only provided equivalent dosages. In terms of adjuvant therapies, 4 studies [93,113,150,160] included tocilizumab as adjuvants, while 1 study [135] used interferon- α 2b as an adjuvant.

Forty-six studies (48.4%) [28,83,84,86,90,91,96,98,100,101,104,107–110,112,115–117,120–122,124,130,132,135,137,139,141,143–145,150,151,154,156–158,164,165,168–170,172,173] included all hospitalized COVID-19 patients, 20 studies (21.1%) [87,93,97,99,102,103,126,131,133,134,136,138,142,146,147,152,155,160,163,174] only included hospitalized patients with severe disease, while 6 studies (6.3%) [88,92,106,111,119,128] included hospitalized patients without severe disease. Last, 22 studies (23.2%) [29,30,85,89,94,105,113,114,118,123,125,127,129,140,148,149,153,159,161,162,167,171] only included ICU hos-

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Figure 2. Risk of bias. (A) Percentage of studies with risk of bias ratings for randomized controlled trials using RoB2. **(B)** Percentage of studies with risk of bias ratings for observational studies using ROBINS-I. ROBINS-I. Risk of bias in non-randomized studies of interventions; RoB2: Revised Cochrane risk of bias tool for randomized trial.

pitalized patients, and 1 study (1.1%) [166] only included hospitalized patients receiving mechanical ventilation. Characteristics of the included studies are tabulated in Supplementary Table 12.

Risk of bias

The risk of bias in the included RCTs was evaluated using RoB2 (Figure 2A & Supplementary Table 10). Seven RCTs (70.0%) [29,30,85,86,88–90] were rated as having some concerns in regards to their risk of bias, mainly due to the use of open-label study designs. The remaining 3 RCTs [28,84,87] were rated as having a low risk of bias. No RCT was rated as having a high risk of bias.

The risk of bias in our included observational studies were assessed using ROBINS-I (Figure 2B & Supplementary Table 11). Thirty-one studies (36.5%) [92,95,96,98,99,103,105–107,109,112,115, 118,120,121,126,133,135,139,143,149,150,152,156,160,165,166,170,172–174] were rated as having critical risk of bias, and 33 studies (38.8%) [91,94,97,100,101,108,110,114,117,119,124,125,128–132,134,136,138,140,141,144–147,158,161,163,164,167–169] were rated as having serious risk of bias. Generally, studies were rated as having a critical or serious risk of bias due to confounding factors and poor descriptions of the regimens/interventions used. Fourteen studies (16.5%) [93,102,104,111,113,116,122,137,142,151,154,155,157,162] were rated as having moderate risk of bias, with no studies receiving a rating of low risk of bias. The remaining 7 studies (8.2%) [83,123,127,148,153,159,171] were rated as 'no information' due to insufficient reporting of study methodology, treatment descriptions and baseline information.

Treatment outcomes

Time to negative test conversion

Nineteen studies [86,102,106,112,119,121,128,129,131–133,136–138,164,165,169,170,172] with 2785 patients evaluated the effect of corticosteroid administration on time to negative test conversion and were included in the meta-analysis (Figure 3). The overall pooled MD was 1.80 (95% CI: 0.48–3.11) with serious and significant heterogeneity ($I^2 = 78\%$; $P_Q < 0.01$). There was no noticeable asymmetry on the funnel plot (Supplementary Figure 1), and Egger's regression test showed no significant small-study effects ($P_{Egger} = 0.82$). There were no significant subgroup

	E	xperin	nental		Contro	bl				Weight
Study (year)	Total	Mean	SD	Total	Mean	SD	Mean difference	MD	95% CI	(%)
Xia et al. 2021 (severe cohort)	46	6.59	3.2200	3	13.00	6.0800 🗲		-6.41	[-13.35; 0.53]	2.3
Ma et al. 2020 B	47	16.10	6.1100	25	19.40	9.4200		-3.30	[-7.38; 0.78]	3.9
Liang et al. 2021	200	25.70	11.9477	200	27.65	11.2009		-1.95	[-4.22; 0.32]	5.4
Chen et al. 2020 B	209	20.32	6.8470	162	21.46	8.9260		-1.14	[-2.80; 0.52]	5.9
Xia et al. 2021 (non-severe cohort)	10	5.90	2.8800	16	6.07	4.7900		-0.17	[-3.12; 2.78]	4.8
Wu et al. 2020 B	226	18.65	6.7149	156	18.70	7.4827		-0.05	[-1.52; 1.41]	6.0
Yuan et al. 2020	35	20.09	7.4206	35	19.76	12.9861		0.33	[-4.62; 5.29]	3.3
You et al. 2020	44	11.89	4.9814	44	11.06	6.8973		0.82	[-1.69; 3.34]	5.2
Zha et al. 2020	11	15.00	1.6965	20	14.00	4.7871		1.00	[-1.33; 3.33]	5.4
Liu et al. 2020 A	72	20.35	8.3215	123	18.30	7.5015	- i	2.06	[-0.28; 4.39]	5.3
Hu et al. 2020 A	28	18.64	7.1867	44	16.29	6.1309		2.35	[-0.87; 5.57]	4.6
Luo et al. 2020	17	12.23	1.5500	15	9.66	1.9200		2.57	[1.35; 3.79]	6.1
Huang et al. 2020	89	18.70	6.0285	220	16.00	4.4774		2.70	[1.32; 4.09]	6.0
Ding et al. 2020	36	11.71	7.7208	46	8.56	6.1220	· · · ·	3.15	[0.07; 6.23]	4.7
Li et al. 2020 B	86	10.88	9.4250	62	7.71	10.6258		3.17	[-0.14; 6.49]	4.5
Ma et al. 2020 A (non-severe cohort)	55	20.71	7.6127	55	17.35	6.8514		3.35	[0.65; 6.06]	5.0
Ma et al. 2020 A (severe cohort)	11	21.39	10.6030	11	17.79	8.3128		- 3.60	[-4.37; 11.56]	1.9
Tang et al. 2021	43	11.00	7.6696	43	7.29	7.6696		3.71	[0.47; 6.95]	4.6
Gong et al. 2020	18	29.11	6.6100	16	24.44	5.2100		4.67	[0.69; 8.65]	4.0
Li et al. 2020 A	55	17.29	6.0902	55	11.71	6.0902		5.59	[3.31; 7.86]	5.4
Li <i>et al.</i> 2021	68	21.39	7.1953	28	13.86	3.1246		7.53	[5.47; 9.60]	5.6
Random effects model	1406			1379			\diamond	1.80	[0.48; 3.11]	100.0
Heterogeneity: $l^2 = 78\%$; p < 0.01						-10	-5 0 5 1	0		

Figure 3. Pooling of mean differences for the outcome of time to negative conversion of SARS-CoV-2 test. The use of corticosteroids was compared with control groups. Heterogeneity was quantified using I^2 statistics. MD <0 indicates beneficial treatment effects of corticosteroids compared with control groups. MD: Mean difference; SD: Standard deviation.

differences between the pooled results from observational studies versus RCTs (p = 0.26; Supplementary Figure 2). Similarly, there were no significant subgroup differences between pooled results from high risk of bias studies compared with low risk of bias studies (p = 0.77; Supplementary Figure 3).

Ten studies [86,106,112,119,121,136,165,169,170,172] assessing 879 patients compared methylprednisolone versus standard of care, while the remaining studies did not disclose the exact corticosteroid regimens used (Supplementary Figure 4). Including only studies that administered methylprednisolone, the pooled MD was 1.81 (95% CI: 0.33–3.29) with substantially decreased heterogeneity compared with the overall analysis (I² = 32%; P_Q = 0.15). There were no significant subgroup differences between subgroups of studies assessing low fixed-dose regimens versus low weight-adjusted dose regimens (p = 0.53; Supplementary Figure 5), nor were there significant subgroup differences between subgroups of studies providing early administration of corticosteroids versus late administration of corticosteroids (p = 0.95; Supplementary Figure 6). Meta-regression analyses did not demonstrate significant correlations between the treatment effect and the proportion of severe patients (p = 0.70; Supplementary Figure 7) or baseline mean SOFA scores (p = 0.53; Supplementary Figure 8).

Three observational studies [120,122,143] reported median time to negative test conversion with IQR, however, these data could not be imputed due to a lack of normality. Based on the Kruskal–Wallis H test, the Mann–Whitney U test and the Student's t-test, all three studies reported no significant differences in the duration to negative test conversion in patients administered corticosteroids versus control.

Length of stay

Thirty studies [86,88,90,92–94,98,102–104,107,115,116,122,126,128,134,135,138,144,147,151,162,164–167,170,172,173] assessing 10,454 patients compared the effect of corticosteroids versus control on length of stay and were included in the meta-analysis (Figure 4). The overall pooled MD was 1.60 (95% CI: -0.20–3.40) with serious and significant heterogeneity ($I^2 = 91\%$; $P_Q < 0.01$). There was no funnel plot asymmetry based on visual inspection (Supplementary Figure 9), and Egger's regression test showed no significant small-study effects ($P_{Egger} = 0.36$).

Subgroup analysis by different study design identified significant between-group differences between the pooled treatment effect from observational studies compared with RCTs (p < 0.05; Supplementary Figure 10). The pooled MD based on 27 observational studies [92–94,98,102–104,107,115,116,122,126,128,134,135,138,144,147,151,162,164–167,170,172,173] was 1.38 (95% CI: -0.62–3.38) with no reduction in heterogeneity (I² = 92%; P_Q < 0.01), while the pooled MD based on 3 RCTs [86,88,90] was 3.62 (95% CI: 1.10–6.15) with no heterogeneity (I² = 0%; P_Q = 0.75).

	E	perime	ental	Control		I				Weight
Study (year)	Total	Mean	SD	Total	Mean	SD	Mean difference	MD	95% CI	(%)
Espinosa-Solano et al. 2021	10	14.98	11.4380	14	27.54	6.5899	←	-12.56	[-20.45; -4.68]	2.1
Badr et al. 2021	32	26.85	20.1769	45	37.35	22.2082	◄	-10.51	[-20.04; -0.97]	1.8
Wang et al. 2020	26	13.64	3.9217	20	22.00	6.3828	←	-8.36	[-11.54; -5.18]	3.2
Aomar-Millán et al. 2021	48	15.65	11.3600	14	22.86	14.5500	← + +	-7.21	[-15.48; 1.06]	2.0
Xia et al. 2021 (severe cohort)	46	17.15	6.3900	3	23.67	8.1400	<	-6.52	[-15.91; 2.87]	1.8
Almas et al. 2021	12	14.23	2.3400	13	20.16	3.4600	——————————————————————————————————————	-5.93	[-8.23; -3.63]	3.4
Cusacovich et al. 2021	117	8.00	11.0000	88	12.00	12.0000		-4.00	[-7.20; -0.80]	3.2
Ma et al. 2020 B	47	18.70	6.7800	25	21.00	7.4900		-2.30	[-5.82; 1.22]	3.2
Ho et al. 2021	574	8.09	6.8000	3739	9.20	24.9000		-1.11	[-2.08; -0.14]	3.6
Xiong et al. 2021 B	47	21.21	9.5700	28	20.79	7.9800		0.42	[-3.61; 4.45]	3.0
He et al. 2021 (severe patients)	40	31.06	9.9961	11	30.57	9.3306		0.49	[-5.83; 6.82]	2.5
Chen et al. 2021	280	16.35	9.6873	148	15.00	7.4865	+	1.35	[-0.31; 3.01]	3.5
Liu et al. 2020 B	409	15.00	8.9274	365	13.00	7.4427		2.00	[0.85; 3.15]	3.5
Batirel et al. 2021 (high dose)	150	12.00	4.4913	150	10.00	5.9884		2.00	[0.80; 3.20]	3.5
Kocakaya et al. 2021	13	18.50	7.8000	12	16.30	4.7000		2.20	[-2.80; 7.20]	2.8
Zha et al. 2020	11	19.63	2.5447	20	17.36	3.1914	- <u></u> -	2.27	[0.22; 4.33]	3.4
Rodríguez-Molinero et al. 2021	25	10.30	6.1224	25	7.50	8.6735		2.80	[-1.36; 6.96]	3.0
Piniella-Ruiz et al. 2021	88	10.26	8.7546	55	7.35	6.8514		2.90	[0.33; 5.48]	3.4
Tang et al. 2021	43	17.35	6.9027	43	14.42	7.6696	· · · ·	2.94	[-0.15; 6.02]	3.3
Ikeda et al. 2021	118	24.35	14.2601	118	21.35	11.2579		3.00	[-0.28; 6.28]	3.2
Corral-Gudino et al. 2021	35	23.00	13.0000	29	20.00	14.0000		- 3.00	[-3.67; 9.67]	2.4
Chen et al. 2020 B	209	17.68	10.1620	162	14.57	7.3620		3.10	[1.32; 4.89]	3.5
Zhan et al. 2021 (severe cohort)	178	14.00	5.9792	90	10.35	3.7671		3.65	[2.47; 4.82]	3.5
Yuan et al. 2020	35	23.86	7.7298	35	19.81	8.7347		4.05	[0.18; 7.91]	3.1
Jamatti et al. 2021	25	11.00	7.8612	25	6.36	3.9306		4.64	[1.20; 8.09]	3.2
Xia et al. 2021 (non-severe cohort)	10	18.60	9.2900	16	13.63	7.4700		+ 4.97	[-1.85; 11.79]	2.4
Wu et al. 2020 B	226	14.70	8.9532	156	9.65	5.2379		5.05	[3.63; 6.48]	3.5
He et al. 2021 (mild/moderate patients)	38	22.36	8.4746	124	16.70	7.5008		5.65	[2.65; 8.65]	3.3
Xiong et al. 2021 A	428	26.35	9.6698	1019	20.35	9.6507		6.00	[4.91; 7.09]	3.5
Li et al. 2020 A	55	22.65	8.3740	55	15.71	6.0902		- 6.94	[4.20; 9.68]	3.3
Obata et al. 2021	57	14.82	6.4648	169	6.00	2.9909		► 8.82	[7.09; 10.56]	3.5
Lu et al. 2021 (no early IFN cohort)	84	26.38	7.1659	87	17.18	7.1616		•• 9.20	[7.06; 11.35]	3.4
He et al. 2021(critical patients)	25	41.00	14.1501	10	31.26	12.0400		▶ 9.74	[0.44; 19.04]	1.8
Random effects model	3541			6913				1.60	[-0.20; 3.40]	100.0
Heterogeneity: $l^2 = 91\%$; p < 0.01							-10 -5 0 5	10		

Figure 4. Pooling of mean differences for the outcome of length of stay. The use of corticosteroids was compared with control groups. Heterogeneity was quantified using I^2 statistics. MD <0 indicates beneficial treatment effects of corticosteroids compared with control groups.

IFN: Interferon; MD: Mean difference; SD: Standard deviation.

There were no significant subgroup differences between pooled results from high risk of bias studies compared with low risk of bias studies (p = 0.13; Supplementary Figure 11).

In addition, there were significant between-group differences between studies using different types of corticosteroids (p < 0.01; Supplementary Figure 12). One study [88] administered dexamethasone and reported the highest MD in length of stay (MD 4.64 [95% CI: 1.20–8.09]). Fourteen studies [86,90,94,104,107,115,126,147,151,162,165,167,170,172] administered methylprednisolone and reported a pooled MD of 0.28 (95% CI: -2.65–3.22). Last, one study [92] administered prednisone and reported the lowest MD of -5.93 (95% CI: -8.23 to -3.63). The remaining studies did not report the exact corticosteroid regimen used in the study.

In terms of subgroup analysis by dosage, there were no significant between-group differences in the pooled MD from studies administering high fixed-dose, high weight-adjusted dose, low fixed-dose, or low weight-adjusted doses of corticosteroids (p = 0.54; Supplementary Figure 13). There were also no significant differences between the pooled MD from studies with early corticosteroids administration versus studies with late corticosteroid administration (p = 0.14; Supplementary Figure 14). Meta-regression analyses did not demonstrate significant correlations between the treatment effect and the proportion of patients with severe disease (p = 0.08; Supplementary Figure 15) or baseline mean SOFA scores (p = 0.28; Supplementary Figure 16).

ICU length of stay

Seven studies [88,98,113,125,140,162,167] assessing 3233 ICU hospitalized patients compared the effect of corticosteroids on ICU length of stay compared with control and were included in the meta-analysis (Figure 5A). The overall pooled MD was -0.01 (95% CI: -3.50–3.47) with serious and significant heterogeneity ($I^2 = 93\%$; $P_Q < 0.01$). We did not draw the funnel plot and conduct Egger's regression test as there were less than 10 included studies.

 (\mathbf{A})

	Experimental			Control							Weight
Study (year)	Total	Mean	SD	Total	Mean	SD		Mean difference	MD	95% CI	(%)
Wang et al. 2020	26	7.64	2.3530	20	14.28	7.9785	-		-6.64	[-10.25; -3.03]	12.4
Batirel et al. 2021 (high dose)	150	4.92	4.3416	150	9.25	5.4644		-	-4.32	[-5.44; -3.21]	15.0
González-Castro et al. 2021 (TCZ cohort)	49	16.12	16.8028	29	19.85	18.7138	-		-3.73	[-12.01; 4.55]	6.8
González-Castro et al. 2021	58	11.35	9.8829	72	10.29	10.5910			1.06	[-2.47; 4.59]	12.5
Ko et al. 2021	187	13.80	13.9759	75	12.60	13.5000			1.20	[-2.45; 4.85]	12.3
Xiong et al. 2021 B	47	15.23	6.7200	28	13.64	6.5800			1.59	[-1.51; 4.69]	13.0
Jamaati et al. 2021	25	7.36	5.5028	25	3.36	2.3584			4.00	[1.65; 6.35]	13.9
Monedero et al. 2021	691	18.90	14.3800	191	14.20	13.0000			4.70	[2.57; 6.83]	14.1
Random effects model	1233			590					-0.01	[-3.50; 3.47]	100.0
Heterogeneity: $I^2 = 93\%$; p < 0.01							-10	-5 0 5	10		

(B)

	Experimen	ntal	Con	trol				Weight
Study (year)	Events To	otal	Events	Total	Odds ratio	OR	95%-CI	(%)
Tortajada et al. 2021	0	38	25	77 -	←	0.02	[0.00; 0.56]	1.8
Espinosa-Solano et al. 2021	6	10	13	14 ·	•	0.12	[0.01; 1.27]	2.9
Salton et al. 2020	15	83	27	90		0.51	[0.25; 1.06]	6.5
Batirel et al. 2021	37 3	300	32	150		0.52	[0.31; 0.87]	6.9
Majmundar <i>et al.</i> 2020	12	60	47	145		0.52	[0.25; 1.07]	6.5
Corral-Gudino et al. 2021	6	35	8	29		0.54	[0.16; 1.80]	5.4
Aomar-Millán et al. 2021	9	48	4	14		0.58	[0.15; 2.26]	5.0
Albani et al. 2021	56 5	559	131	844		0.61	[0.43; 0.85]	7.2
Cusacovich et al. 2021	19 1	117	21	88		0.62	[0.31; 1.24]	6.6
Li et al. 2021	45	68	21	28		0.65	[0.24; 1.76]	5.9
Bartoletti et al. 2020	32 1	170	84	343		0.71	[0.45; 1.13]	7.0
Ho et al. 2021	39 5	574	285	3739		0.88	[0.62; 1.25]	7.2
Tang et al. 2021	2	43	2	43		1.00	[0.13; 7.44]	3.6
Fusina et al. 2021	53 4	480	56	831		1.72	[1.16; 2.55]	7.1
Liu et al. 2020 B	195 4	409	82	365		3.14	[2.30; 4.30]	7.3
Obata et al. 2021	17	57	16	169		4.06	[1.89; 8.74]	6.4
Chen et al. 2020 B	53 2	209	11	162		4.66	[2.35; 9.27]	6.6
Random effects model Heterogeneity: I ² = 87%; p < 0.01	32	3260		7131	0.1 0.2 0.5 1 2 5 10	0.87	[0.52; 1.47]	100.0
Liu et al. 2020 B Obata et al. 2021 Chen et al. 2020 B Random effects model Heterogeneity: $I^2 = 87\%$; p < 0.01	195 4 17 53 2 32	409 57 209 3260	82 16 11	365 169 162 7131		3.14 4.06 4.66 0.87	[2.30; 4.30] [1.89; 8.74] [2.35; 9.27] [0.52; 1.47]	- 6 6 100

Figure 5. Pooling of mean differences for intensive care unit length of stay and odds ratio for incidence of intensive care unit admission. The use of corticosteroids was compared with control groups. Heterogeneity was quantified using I^2 statistics. (A) Forest plot for the pooling of MDs for ICU length of stay. MD <0 indicates beneficial treatment effects of corticosteroids compared with control groups. (B) Forest plot for the pooling of ORs for incidence of ICU admission. OR <1 indicates beneficial treatment effects of corticosteroids compared with control groups. MD: Mean difference; OR: Odds ratio; SD: Standard deviation; TCZ: Tocilizumab.

Subgroup analysis by study design revealed significant differences between the overall MD from observational studies versus RCTs (p < 0.05; Supplementary Figure 17). Across the 6 observational studies [98,113,125,140,162,167], the pooled MD was -0.66 (95% CI: -4.48–3.16) with no reduction in heterogeneity ($I^2 = 92\%$; $P_Q < 0.01$). Only one study [88] was included in the RCT subgroup, which reported a MD of 4.00 (95% CI: 1.65–6.35). There were no significant subgroup differences between pooled results from high risk of bias studies compared with low risk of bias studies (p = 0.15; Supplementary Figure 18).

There were no significant differences between subgroup results from the subgroup analyses by corticosteroid type (p = 0.14; Supplementary Figure 19) or dose (high fixed-dose vs low weight-adjusted dose, p = 0.75; Supplementary Figure 20). However, there were significant subgroup differences between studies with early administration of corticosteroids versus studies with late administration of corticosteroids (p < 0.01; Supplementary Figure 21). Two studies administering early steroids yielded a pooled MD of 2.00 (-1.63–5.63) with a substantial reduction in heterogeneity compared with the overall meta-analysis (I² = 0%; P_Q = 0.75), while one study administering late corticosteroid yielded an MD of 10.50 (95% CI: 7.57–13.43). Last, meta-regression analyses did not reveal any significant correlations between the treatment effect and the proportion of patients with severe disease (p = 0.82; Supplementary Figure 22) or baseline mean SOFA scores (p = 0.87; Supplementary Figure 23).

Five studies [94,105,114,129,134] reported median length of ICU stay with IQR. However, as the data were significantly skewed, they could not be imputed and thus were not included in the meta-analysis. Based on the Mann–Whitney U test, two studies [94,129] did not find a significant difference between the median length of ICU stay in patients taking corticosteroids versus control, two studies [105,134] found that the length of ICU stay is significantly longer in patients taking corticosteroids versus control, and one study [114] found that the length of ICU stay is significantly longer in control patients versus patients taking corticosteroids.

ICU admission

Seventeen studies [86,90,91,93,97,98,102,104,107,110,116,131,134,139,144,155,157] with 10,391 patients compared the effect of corticosteroids versus control on incidence of ICU admission and were included in the meta-analysis (Figure 5B). The overall pooled OR was 0.87 (95% CI: 0.52–1.47) with serious and significant heterogeneity ($I^2 = 87\%$; $P_Q < 0.01$). Visual examination of the funnel plot (Supplementary Figure 24) and Egger's regression test showed no evidence of small-study effects ($P_{Egger} = 0.24$). There were no significant between-group differences in the subgroup analysis by study design (p = 0.42; Supplementary Figure 25) and by risk of bias (p = 0.61; Supplementary Figure 26).

Five studies [86,90,104,107,155] with 552 patients assessed the efficacy of methylprednisolone versus standard of care. The remaining studies did not disclose the exact type of corticosteroid used. The pooled OR for studies using methylprednisolone was 0.53 (95% CI: 0.29–0.97) with no heterogeneity ($I^2 = 0\%$; $P_Q = 0.71$; Supplementary Figure 27).

There were no significant differences between the subgroup results from studies using different dosages of corticosteroids (p = 0.36; Supplementary Figure 28), nor were there any differences between the subgroup results from studies administering early corticosteroids versus late corticosteroids (p = 0.43; Supplementary Figure 29). Last, meta-regression analyses did not reveal significant correlations between the treatment effect and the proportion of patients with severe disease (p = 0.52; Supplementary Figure 30) and baseline mean SOFA score (p = 0.29; Supplementary Figure 31).

Mortality

Eighty-six studies [28–30,83–94,96–99,101–108,110–125,127–129,131–135,137–142,144–164,166–169,171–174] with 40,623 patients compared the effect of corticosteroids versus control in terms of mortality incidence and were included in the meta-analysis (Figure 6). The overall pooled OR was 1.00 (95% CI: 0.84–1.20) with serious and significant heterogeneity ($I^2 = 76\%$; $P_Q < 0.01$). Visual inspection of the funnel plot (Supplementary Figure 32) and Egger's regression test revealed no evidence of small-study effects ($P_{Egger} = 0.44$). Subgroup analyses showed that there were no significant differences in the pooled OR between studies of different designs (p = 0.27; Supplementary Figure 33) or between studies rated as high risk of bias versus studies with a low risk of bias (p = 0.17; Supplementary Figure 34).

Subgroup analyses by different corticosteroid regimens revealed significant between-group differences (p < 0.05; Supplementary Figure 35). Eight studies [28,30,88,105,114,117,141,171] using dexamethasone yielded a pooled OR of 0.94 (95% CI: 0.78–1.13) with no heterogeneity (I² = 0%; P_Q = 0.72). Three studies [29,85,87] assessed the efficacy of hydrocortisone and yielded a pooled OR of 0.88 (95% CI: 0.11–6.84) with moderate heterogeneity (I² = 55%; P_Q = 0.11). Thirty studies [84,86,89,90,94,99,101,104,106,107,112,115,119,121,138,142,146,147,151–155,161,162,167–169,172,174] assessed the efficacy of methylprednisolone and yielded a pooled OR of 0.62 (95% CI: 0.45–0.85) with moderate heterogeneity (I² = 36%; P_Q < 0.05). Last, one study [145] assessed the efficacy of prednisolone with an OR of 0.38 (95% CI: 0.01–13.17), and another [92] assessed the efficacy of prednisone with an OR of 0.06 (95% CI: 0.01–0.59).

There were no significant differences between the subgroup results from studies using different dosages of corticosteroids (p = 0.06; Supplementary Figure 36), nor were there any differences between the subgroup results from studies administering early corticosteroids versus late corticosteroids (p = 0.28; Supplementary Figure 37). Last, meta-regression analyses did not reveal significant correlations between the treatment effect and the proportion of patients with severe disease (p = 0.56; Supplementary Figure 38), baseline mean SOFA score (p = 0.48; Supplementary Figure 39), and follow-up duration (p = 0.14; Supplementary Figure 40).

Incidence of mechanical ventilation

Thirty-seven studies [28,86,88,89,95–98,100–105,107,113,116,122,124,130,131,134,136,138,139,141,144–146,148,155–158,162,164,167] with 20,453 patients compared the effect of corticosteroids versus control in terms of incidence of mechanical ventilation and were included in the meta-analysis (Figure 7A). The overall pooled OR was 1.19 (95% CI: 0.81–1.76) with serious and significant heterogeneity ($I^2 = 84\%$; $P_Q < 0.01$). Visual inspection of the funnel plot (Supplementary Figure 41) and Egger's regression test showed no evidence of small-study effects ($P_{Egger} = 0.90$), and there were no significant between-group differences among studies using different study designs (p = 0.73; Supple-

	Experi	nental	Con	trol				Weight
Study (year)	Events	Total	Events	Total	Odds ratio	OR	95% CI	(%)
Almas et al. 2021	1	12	8	13	←	0.06	[0.01; 0.59]	0.5
Edalatifard et al. 2020	2	34	12	28	→	0.08	[0.02; 0.42]	0.8
Espinosa-Solano et al. 2021	0	10	3	14	•	0.13	[0.00; 3.56]	0.3
Wang & Guerra 2021 Zhu et al. 2020	40	18	20	11		0.18	[0.03; 1.07]	0.7
Saggi et al. 2020	10	37	24	38	*	0.22	[0.08; 0.58]	1.2
Salton et al. 2020	6	83	21	90	▲	0.26	[0.10; 0.67]	1.2
Tang et al. 2021	0	43	1	43	→	0.33	[0.01; 8.22]	0.3
You et al. 2020 González-Castro et al. 2021	10	44 58	27	44 72	\leftarrow	0.33	[0.01; 8.22]	0.3
Russo et al. 2020	14	63	30	67		0.35	[0.16; 0.76]	1.4
Ko et al. 2021	39	187	31	75		0.37	[0.21; 0.67]	1.5
Ooi et al. 2020	0	35	1	57	< ■	0.38	[0.01; 13.17]	0.3
González-Castro et al. 2021 (TCZ Cohort)	7	49	8	29	· · · · · · · · · · · · · · · · · · ·	0.44	[0.14; 1.37]	1.1
Piniella-Buiz et al. 2021	60	88	20 45	55		0.45	[0.20, 1.02]	1.3
Majmundar et al. 2020	8	60	34	145		0.50	[0.22; 1.16]	1.3
Ma et al. 2020 B	2	47	2	25	< <u>-</u>	0.51	[0.07; 3.87]	0.6
Fernández-Cruz et al. 2020	55	396	16	67		0.51	[0.27; 0.97]	1.5
Lambermont et al. 2020	33 20	58	38 91	88 189		0.52	[0.29; 0.93]	1.5
Van den Eynde et al. 2021	17	38	69	118		0.57	[0.28; 1.20]	1.4
Batirel et al. 2021	30	300	24	150		0.58	[0.33; 1.04]	1.5
Borie et al. 2020	22	108	19	63		0.59	[0.29; 1.21]	1.4
Hoertel et al. 2021 (respiratory support cohort)	10	122	150	196		0.60	[0.30; 1.22]	1.4
Rahman et al. 2021	15	72	19	64		0.62	[0.29: 1.36]	1.4
Jung et al. 2021	406	967	1120	2115	-	0.64	[0.55; 0.75]	1.7
Xue et al. 2020	2	60	3	60		0.66	[0.11; 4.07]	0.7
Nelson et al. 2021	13	42	17	42		0.66	[0.27; 1.62]	1.3
Gallay et al. 2021	43	98 151	84 Q1	148		0.76	[0.46; 1.26]	1.5
Hong et al. 2020	49	93	48	84		0.84	[0.46; 1.51]	1.5
Van den Eynde et al. 2021 (TCZ adjuvant)	23	78	7	21		0.84	[0.30; 2.34]	1.2
Ho et al. 2021	153	574	1117	3739		0.85	[0.70; 1.04]	1.7
RECOVERY 2021 REMAR-CAP 2020	482	2104	1110	4321		0.86	[0.76; 0.97]	1.7
Keller et al. 2020	19	140	251	1666		0.89	[0.54: 1.46]	1.5
Monedero et al. 2021	238	691	70	191		0.91	[0.65; 1.27]	1.6
Jeronimo et al. 2021	79	209	80	207	- <u>+</u> -	0.96	[0.65; 1.43]	1.6
Ruiz-Irastorza et al. 2020	6	61	18	181		0.99	[0.37; 2.61]	1.2
Xiong et al. 2021 A Gong et al. 2020	0	428	0	1019		1.00	[0.01; 73.40]	0.2
Zha et al. 2020	0	11	0	20	← →	1.00	[0.02; 64.06]	0.2
Rodríguez-Molinero et al. 2021	5	25	5	25		1.00	[0.25; 4.00]	0.9
Hu et al. 2020 B	0	216	0	92	→ →	1.00	[0.01; 72.91]	0.2
Hu et al. 2020 A He et al. 2021 (mild/moderate patients)	0	28	0	44 124		1.00	[0.02; 57.28]	0.2
Liu et al. 2020 A	0	72	0	123	<→	1.00	[0.02; 58.65]	0.2
Huang et al. 2020	0	89	0	220	<→	1.00	[0.01; 76.33]	0.2
Ritter et al. 2021	22	74	18	61		1.01	[0.48; 2.12]	1.4
Zhan et al. 2021	15	257	16	286		1.05	[0.51; 2.16]	1.4
Hoertel et al. 2021	14	108	131	1080		1.00	[0.60: 1.95]	1.5
Badr et al. 2021	10	32	13	45		1.12	[0.42; 3.00]	1.2
Zawaydeh et al. 2021	12	41	24	89		1.12	[0.49; 2.54]	1.3
Papamanoli et al. 2021	37	153	65	294		1.12	[0.71; 1.78]	1.6
Corral-Gudino et al. 2021	7	25	5	25		1.19	[0.34; 4.28]	1.0
Liang et al. 2021	6	200	5	200		1.21	[0.36; 4.02]	1.0
Bartoletti et al. 2020	60	170	106	343		1.22	[0.83; 1.80]	1.6
Tran et al. 2021	17	203	46	688		1.28	[0.71; 2.28]	1.5
Pusina et al. 2021 De Greef et al. 2020	125	480	1/9	29		1.28	[0.99; 1.67]	1.7
Mühlemann et al. 2021	18	101	13	93		1.33	[0.61; 2.90]	1.4
Rodríguez-Baño et al. 2021	30	195	41	344		1.34	[0.81; 2.23]	1.5
Gragueb-Chatti et al. 2021	28	84	18	67		1.36	[0.67; 2.76]	1.4
Lietal 2020 B	97	55 183	49	95		1.42	[0.63; 3.20]	1.3
Aomar-Millán <i>et al.</i> 2021	5	48	1	14	→	1.51	[0.16; 14.13]	0.5
Wang et al. 2020	2	26	1	20	← → →	1.58	[0.13; 18.81]	0.5
Tortajada et al. 2021	16	38	23	77		1.71	[0.76; 3.83]	1.3
Liu et al. 2020 B	185	409	115	365		1.80	[1.34; 2.41]	1.7
Spagnuolo et al. 2020	4	59	8	221		1.92	[0.56: 6.67]	1.0
Balena et al. 2020	17	42	34	148		2.28	[1.10; 4.71]	1.4
He et al. 2021 (severe patients)	1	40	0	11		2.33	[0.03; 213.51]	0.2
Chen et al. 2020 B	102	209	43	162		2.64	[1.70; 4.10]	1.6
lkeda et al. 2021	14	118	5	118		3.04	[1.06: 8.74]	1.2
Li et al. 2020 A	1	55	0	55		3.06	[0.12; 76.64]	0.3
He et al. 2021 (critical patients)	17	25	4	10		3.19	[0.70; 14.56]	0.9
Ding et al. 2020	1	36	0	46	* * * *	3.37	[0.15; 76.00]	0.3
Chen et al. 2020 A	2	64	0	87 49	*	4.58 4.68	[0.18; 117.58] [0.18: 119.351	0.3
Chen et al. 2021	151	280	29	148		4.80	[3.01; 7.68]	1.6
Obata et al. 2021	23	55	18	169		6.03	[2.92; 12.45]	1.4
Rodríguez-Baño et al. 2021 (TCZ adjuvant)	19	151	2	88		6.19	[1.41; 27.25]	0.9
wu eral. 2020 A Lietal. 2021	83	531 69	26	983	→ 	6.82 7.67	[4.33; 10.74] [2 72: 21 60]	1.6 1.2
Xiong et al. 2021 B	4	47	0	28		8.06	[0.28; 233.90]	0.3
Ma et al. 2020 A	3	126	0	324		12.00	[0.93; 155.10]	0.4
Lu et al. 2021 (no early IFN cohort)	11	84	0	87		26.89	[1.59; 453.49]	0.4
Pandom offects model		14 460		26 15/	\downarrow	1.00	[0 94- 1 00]	100.0
Heterogeneity: $l^2 = 76\%$; p < 0.01		. 4,403		_0,104		1.00	[0.04; 1.20]	100.0
					0.2 0.5 1 2 5			

Figure 6. Pooling of odds ratios for the outcome of mortality incidence. The use of corticosteroids was compared with control groups. Heterogeneity was quantified using I^2 statistics. OR <1 indicates beneficial treatment effects of corticosteroids compared with control groups.

IFN: Interferon; OR: Odds ratio; TCZ: Tocilizumab.

A	Experin	nental	Con	trol				M
Study (year)	Events	Total	Events	Total	Odds ratio	OR	95% CI	(%)
	Evento	10101	LVCIILO	Total		0.00	10.00.0.051	(70)
Espinosa-Solano et al. 2021	2	10	13	14	• :	0.02	[0.00; 0.25]	1.5
Tortajada et al. 2021	0	38	4	//	▲■	0.14	[0.00; 4.73]	1.0
Luo et al. 2020	0	17	2	15	<	0.17	[0.01; 3.48]	1.2
Gonzalez-Castro et al. 2021 (ICZ cohor	1) 42	49	28	29	<	0.21	[0.02; 1.84]	1.8
Li et al. 2020 C	5	47	7	21	<	0.24	[0.07; 0.87]	2.6
Wang et al. 2020	1	26	2	20	<	0.36	[0.03; 4.28]	1.5
Mühlemann <i>et al.</i> 2021	38	101	54	93		0.44	[0.24; 0.77]	3.3
Batirel et al. 2021	29	300	28	150		0.47	[0.27; 0.82]	3.3
Salton et al. 2020	14	83	26	90		0.50	[0.24; 1.04]	3.1
Cusacovich et al. 2021	16	117	21	88		0.51	[0.25; 1.04]	3.1
Keller et al. 2020	7	140	128	1666		0.63	[0.29; 1.38]	3.1
Majmundar <i>et al.</i> 2020	11	60	36	145		0.68	[0.32; 1.45]	3.1
Tran et al. 2021	19	203	85	688		0.73	[0.43; 1.24]	3.3
RECOVERY 2021	110	1780	298	3638		0.74	[0.59; 0.93]	3.5
Papamanoli <i>et al.</i> 2021	50	153	114	294		0.77	[0.51; 1.16]	3.4
Bartoletti et al. 2020	30	170	73	343		0.79	[0.49; 1.27]	3.3
Ma et al. 2020 B	5	47	3	25		0.87	[0.19; 4.00]	2.3
Ooi et al. 2020	4	35	7	57	,	0.92	[0.25: 3.41]	2.6
Edalatifard et al. 2020	0	34	0	28	<→	1.00	[0.02: 52.96]	0.8
Spagnuolo et al. 2020	21	59	75	221		1.08	[0.59: 1.96]	3.2
Ho et al. 2021	86	574	495	3739		1.15	[0.90: 1.48]	3.4
Balena et al. 2020	5	42	14	148		1.29	[0.44: 3.82]	2.8
Jamaati et al. 2021	13	25	11	25		1.38	[0.45: 4.20]	2.8
Rahman et al. 2021	56	72	41	64		1.96	[0.92: 4.18]	3.1
Liu et al. 2020 B	84	409	42	365		1.99	[1.33: 2.97]	3.4
Li et al. 2021	36	68	10	28		2.02	[0.82: 5.02]	3.0
Chen et al. 2020 B	20	209	8	162		2.04	[0.87: 4.75]	3.0
De Greef et al. 2020	13	18	16	29		2.11	[0.60: 7.48]	2.6
Ikeda et al. 2021	47	118	28	118	· · · · · · · · · · · · · · · · · · ·	2 13	[1 21: 3 73]	3.3
González-Castro et al. 2021	48	58	48	72		2 40	[1.04:5.55]	3.0
Wulet al. 2020 B	167	226	84	156		2 43	[1.57: 3.74]	3.4
Chen et al. 2021	130	280	30	148		3 41	[2 14: 5 42]	3.3
Babl et al. 2021	243	846	65	615		3.41	[2 53: 4 59]	3.4
Tang et al. 2021	240	13	1	/3		/ 31	[0.46:40.23]	17
Chaudhry et al. 2021	2	90	0	72		4.56	[0.17.110.32]	1.1
Chan at $a/2020$ A	17	64	1	10		17.36	[0.17, 113.02]	1.1
Viena et al. 2021 P	10	47	0	49		01.46	[2.22, 100.74]	1.0
Along El al. 2021 D	20	47 55	6	160		20.20		2.0
	29	55	0	109		50.50	[11.47, 00.00]	2.9
Random effects model		6721	1	3,732	÷	1.19	[0.81; 1.76]	100.0
Heterogeneity: $l^2 = 84\%$: p < 0.01		-	-				1 , .	
				(0.1 0.2 0.5 1 2 5 10			

B

	Ex	perime	ental		Contr	ol				Weight
Study (year)	Total	Mean	SD	Total	Mean	SD	Mean difference	MD	95% CI	(%)
Mühlemann et al. 2021	38	32.26	35.4393	54	39.17	30.4659		-6.91	[-20.80; 6.98]	0.7
Badr et al. 2021	32	14.71	13.1926	45	18.42	13.7844		-3.71	[-9.80; 2.39]	3.5
Ma et al. 2020 B	26	9.60	6.3600	8	12.80	6.4000		-3.20	[-8.26; 1.86]	5.0
Urendes et al. 2020	55	10.70	8.2000	95	13.40	10.3000		-2.70	[-5.70; 0.30]	12.0
Kocakaya <i>et al.</i> 2021	13	6.23	3.4000	12	7.67	2.1000		-1.44	[-3.64; 0.76]	18.7
Tomazini <i>et al.</i> 2020	151	12.50	8.1503	148	13.90	7.4483		-1.40	[-3.17; 0.37]	24.2
Batirel et al. 2021 (low dose)	150	6.11	4.1170	150	7.32	5.6141		-1.21	[-2.32; -0.10]	35.9
Random effects model Heterogeneity: $I^2 = 0\%$; $p = 0$	465 .88			512		-	→ 10 -5 0 5	-1.71	[–2.54; –0.88]	100.0

Figure 7. Pooling of odds ratios for incidence of mechanical ventilation and mean differences for length of mechanical ventilation. The use of corticosteroids was compared with control groups. Heterogeneity was quantified using l^2 statistics. (A) Forest plot for the pooling of ORs for incidence of mechanical ventilation. OR <1 indicates beneficial treatment effects of corticosteroids compared with control groups. (B) Forest plot for the pooling of MDs for length of mechanical ventilation. MD <0 indicates beneficial treatment effects of corticosteroids compared with control groups.

MD: Mean difference; OR: Odds ratio; SD: Standard deviation; TCZ: Tocilizumab.

mentary Figure 42) or studies with high risks of bias versus studies with low risks of bias (p = 0.45; Supplementary Figure 43).

There were no significant differences between studies using different corticosteroid regimens (p = 0.95; Supplementary Figure 44), studies using different corticosteroid doses (p = 0.56; Supplementary Figure 45), or between studies administering early corticosteroids versus late corticosteroids (p = 0.89; Supplementary Figure 46). Last, meta-regression analyses revealed no significant correlations between the treatment effect and proportion of patients with severe diseases (p = 0.63; Supplementary Figure 47) and baseline mean SOFA score (p = 0.88; Supplementary Figure 48).

Length of mechanical ventilation

Seven studies [30,94,98,126,138,141,159] with 977 patients receiving mechanical ventilation compared the effect of corticosteroids versus control on the length of mechanical ventilation and were included in the meta-analysis (Figure 7B). Overall, the pooled MD was -1.71 (95% CI: -2.54 to -0.88) with no heterogeneity ($I^2 = 0\%$; $P_Q = 0.88$). Because there were less than 10 included studies, we did not draw funnel plots nor did we conduct Egger's regression test. The subgroup analysis by study design (Supplementary Figure 49) and risk of bias (Supplementary Figure 50) showed no significant between-group differences (p = 0.67 for both subgroup analyses).

The subgroup analysis by different corticosteroid regimens (Supplementary Figure 51) and corticosteroid dosage (Supplementary Figure 52) also did not show significant between-group differences (p = 0.91 and p = 0.99 for regimen and dosage subgroups, respectively). Last, meta-regression analyses by the proportion of severe patients (Supplementary Figure 53) and baseline mean SOFA score (Supplementary Figure 54) did not reveal any significant correlations (p = 0.30 and p = 0.39 for the severe proportions and SOFA score regressions, respectively). We did not conduct subgroup analyses by time to administration due to a lack of data.

Five studies [87,114,122,129,155] reported median length of mechanical ventilation with IQR, however, they were not imputed and included in the meta-analysis due to skewed data. One study [87] did not perform any statistical testing, three studies [114,129,155] used the Mann–Whitney U test, and one study [122] used Student's t-test. Out of the four studies with statistical tests, one study [122] (using the t-test) did not find significant differences between the corticosteroid and control groups. Two studies [114,155] found a significantly shorter median length of mechanical ventilation in patients taking corticosteroids, and one study [129] found a significantly shorter median length of mechanical ventilation of the control patients. The study that did not conduct statistical tests [87] reported a median duration of mechanical ventilation of 15 h (IQR: 13–21 h) in the control group, compared with a median duration of 6 h (IQR: 4–20 h) in the corticosteroids group.

Adverse events

Nine studies [30,96,102,111,133,140,155,158,172] with 3409 patients reported incidence of adverse events in the corticosteroid and control groups (Figure 8A). The overall pooled OR was 1.89 (95% CI: 0.98–3.63) with moderate and significant heterogeneity ($I^2 = 71\%$; $P_Q < 0.01$). There were significant differences between the treatment effect from observational studies versus RCTs (p < 0.01; Supplementary Figure 55), where 8 observational studies [96,102,111,133,140,155,158,172] yielded a pooled OR of 2.17 (95% CI: 1.24–3.81) with similar heterogeneity compared with the overall analysis, and 1 RCT [30] reported an OR of 0.43 (95% CI: 0.14–1.26). In terms of risk of bias, there were no significant between-group differences in the subgroup analysis (p = 0.46; Supplementary Figure 56). We did not draw funnel plots nor conducted Egger's regression tests because less than 10 studies were included in the analysis.

There were no significant between-group differences by corticosteroid regimen (p = 0.08; Supplementary Figure 57), corticosteroid dosage (p = 0.05; Supplementary Figure 58) and time to administration (p = 0.79; Supplementary Figure 59). Meta-regression analysis showed no significant correlation associated with baseline mean SOFA score (p = 0.74; Supplementary Figure 60). We did not conduct meta-regression analysis by the proportion of patients with severe disease due to a lack of data.

Severe adverse events

Six studies [29,30,87,89,104,120] with 1142 patients reported incidence of severe adverse events (Figure 8B). The pooled OR was 0.87 (95% CI: 0.47–1.62) with no heterogeneity ($I^2 = 0\%$; $P_Q = 0.77$). There were no significant between-group differences in the subgroup analysis by study design (p = 0.74; Supplementary Figure 61) or risk of

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A	Experin	nental	Con	trol				Weight
Study (year)	Events	Total	Events	Total	OR	OR	95% CI	(%)
Tomazini <i>et al.</i> 2020	5	151	11	148		0.43	[0.14; 1.26]	10.2
Salton <i>et al.</i> 2020	29	83	30	90		1.07	[0.57; 2.01]	13.0
Tran <i>et al.</i> 2021	150	283	314	682		1.32	[1.00; 1.75]	14.7
Zha <i>et al.</i> 2020	5	11	7	20		1.55	[0.34; 6.94]	7.9
Balena <i>et al.</i> 2020	13	42	32	148		1.62	[0.76; 3.48]	12.2
Chen et al. 2020 B	109	209	44	162		2.92	[1.88; 4.54]	14.0
Gallay et al. 2021	22	131	11	172		2.95	[1.38; 6.34]	12.2
Monedero et al. 2021	681	691	179	191		4.57	[1.94; 10.74]	11.7
Liu et al. 2020 A	6	72	0	123		18.81	[1.39; 254.33]	4.0
Random effects model		1673		1736	\sim	1.89	[0.98; 3.63]	100.0
Heterogeneity: I ² = 71%; p < 0.01				0.	1 0.2 0.5 1 2 5 10			

B	Experim	nental	Con	trol				Weight
Study (year)	Events	Total	Events	Total	OR	OR	95% CI	(%)
Tomazini <i>et al.</i> 2020	5	151	9	148		0.53	[0.17; 1.62]	28.1
Edalatifard et al. 2020	2	34	2	28		0.81	[0.11; 6.17]	10.9
Cusacovich et al. 2021	17	117	15	88		0.83	[0.39; 1.76]	44.5
Hu et al. 2020 B	0	216	0	92	• • • • •	1.00	[0.01; 72.91]	2.7
REMAP-CAP 2020	4	137	1	101		3.01	[0.33; 27.32]	9.4
Munch et al. 2021	1	16	0	14		3.06	[0.11; 87.58]	4.3
Random effects model Heterogeneity: I ² = 0%; p = 0.77		671		471 0	.1 0.2 0.5 1 2 5 10	0.87	[0.47; 1.62]	100.0

C	Experin	nental	Con	trol				Weiaht
Study (year)	Events	Total	Events	Total	OR	OR	95% CI	(%)
Tang et al. 2021	3	43	7	43	→ ■ ↓ ↓	0.39	[0.09; 1.60]	10.1
Tomazini et al. 2020	47	151	42	148		1.14	[0.69; 1.87]	14.3
Wu et al. 2020 B	20	226	12	156		1.17	[0.55; 2.46]	13.3
Cusacovich et al. 2021	10	117	5	88		1.55	[0.51; 4.71]	11.6
Tran et al. 2021	64	283	86	682		2.03	[1.41; 2.90]	14.7
Luo et al. 2020	1	17	0	15		3.06	[0.11; 86.63]	4.1
Hu et al. 2020 B	19	216	2	92	• •	4.34	[0.99; 19.03]	9.9
Xue et al. 2020	4	60	0	60		9.64	[0.51; 183.05]	4.9
Gallay et al. 2021	8	98	1	166		14.67	[1.81; 119.13]	7.3
Salton et al. 2020	8	83	0	90		19.55	[1.17; 326.68]	5.2
Corral-Gudino et al. 2021	9	35	0	29		23.38	[1.13; 481.76]	4.7
Random effects model		1329		1569		2.45	[1.08; 5.56]	100.0
Heterogeneity: $I^2 = 53\%$; $p = 0.02$				0	1 0.2 0.5 1 2 5 10			

D	Experin	nental	Con	trol				Weight
Study (year)	Events	Total	Events	Total	I OR	OR	95% CI	(%)
Li et al. 2020 C	1	47	3	21		0.13	[0.01; 1.34]	1.3
González-Castro et al. 2021 (TCZ cohort)	17	49	18	29		0.32	[0.13; 0.84]	3.7
Ritter et al. 2021	36	74	45	61		0.34	[0.16; 0.70]	4.5
Obata et al. 2021	1	55	7	169	< <u>∎</u>	0.43	[0.05; 3.56]	1.5
Espinosa-Solano et al. 2021	4	10	7	14		0.67	[0.13; 3.45]	2.1
Tomazini et al. 2020	33	151	43	148		0.68	[0.40; 1.15]	5.2
Cusacovich et al. 2021	26	117	25	88		0.72	[0.38; 1.36]	4.8
González-Castro et al. 2021	14	58	20	72		0.83	[0.37; 1.83]	4.2
Ho et al. 2021	5	61	279	2890		0.84	[0.33; 2.10]	3.8
Dequin et al. 2020	28	76	30	73		0.84	[0.43; 1.62]	4.7
Rodríguez-Baño et al. 2021	18	190	36	339		0.88	[0.49; 1.60]	4.9
Tran et al. 2021	50	283	128	682		0.93	[0.65; 1.33]	5.7
Rodríguez-Baño et al. 2021 (TCZ adjuvant)	18	150	11	88		0.95	[0.43; 2.13]	4.2
Spagnuolo et al. 2020	6	59	23	221		0.97	[0.38; 2.52]	3.7
Yuan et al. 2020	0	35	0	35	·	1.00	[0.02; 51.80]	0.5
Gragueb-Chatti et al. 2021	56	84	44	67		1.05	[0.53; 2.06]	4.6
Salton et al. 2020	1	83	1	90	· • · · · · · · · · · · · · · · · · · ·	1.09	[0.07; 17.64]	0.9
Liu et al. 2020 B	31	224	17	145		1.21	[0.64; 2.28]	4.8
Nelson et al. 2021	22	42	19	42		1.33	[0.56; 3.14]	4.0
Xue et al. 2020	8	60	6	60		1.38	[0.45; 4.26]	3.2
Russo et al. 2020	30	63	25	67		1.53	[0.76; 3.07]	4.6
Balena et al. 2020	13	42	32	148		1.62	[0.76; 3.48]	4.4
Monedero et al. 2021	417	691	92	191		1.64	[1.19; 2.26]	5.8
Li et al. 2020 B	34	183	10	111		2.30	[1.09; 4.87]	4.4
Edalatifard et al. 2020	1	34	0	28		2.91	[0.10; 83.05]	0.7
Hu et al. 2020 B	21	216	3	92		3.19	[0.93; 10.99]	2.9
Corral-Gudino et al. 2021	5	35	1	29		4.67	[0.51; 42.45]	1.4
De Greef et al. 2020	10	18	5	29		6.00	[1.57; 22.89]	2.7
Gallay et al. 2021	2	98	0	166		6.52	[0.40; 106.31]	0.9
Random effects model		3288		6195	·	1.06	[0.82; 1.37]	100.0
Heterogeneity: $I^2 = 47\%$; p < 0.01				0	0.1 0.2 0.5 1 2 5 10)		

Figure 8. Pooling of odds ratios for all safety outcomes. The use of corticosteroids was compared with control groups. Heterogeneity was quantified using I^2 statistics. **(A)** Forest plot for the pooling of ORs for incidence of adverse events. **(B)** Forest plot for the pooling of ORs for incidence of serious adverse events. **(C)** Forest plot for the pooling of ORs for incidence of hyperglycemia. **(D)** Forest plot for the pooling of ORs for incidence of nosocomial infections. OR <1 indicates beneficial treatment effects of corticosteroids compared with control groups for all safety outcomes. OR: Odds ratio; TCZ: Tocilizumab.

bias (p = 0.96; Supplementary Figure 62). We did not draw funnel plots or conduct Egger's regression test as less than 10 studies were included.

Subgroup analyses by corticosteroid regimen (Supplementary Figure 63) and corticosteroid dosage (Supplementary Figure 64) revealed significant between-group differences (p < 0.01 for both subgroup analyses). Two studies [89,104] used high fixed-dose methylprednisolone and yielded an OR of 0.83 (95% CI: 0.77–0.89) with no heterogeneity ($I^2 = 0\%$; $P_Q = 0.99$), while two studies used low fixed-dose hydrocortisone [29,87] and yielded an OR of 3.02 (95% CI: 2.73–3.34) with no heterogeneity ($I^2 = 0\%$; $P_Q = 0.99$). One study [120] used a low weight-adjusted dose of unspecified corticosteroid and reported an OR of 1.00 (95% CI: 0.01–72.91), and one study [30] compared dexamethasone versus standard of care and reported an OR of 0.53 (95% CI: 0.17–1.62). There were no between-group differences for subgroup analysis by time to administration (p = 0.73; Supplementary Figure 65), and we did not perform any meta-regression analysis due to a lack of data.

Hyperglycemia

Eleven studies [30,86,90,104,111,120,136,155,158,164,168] with 2898 patients reported incidence of hyperglycemia (Figure 8C) with a pooled OR of 2.45 (95% CI: 1.08–5.56). There was significant and moderate heterogeneity ($I^2 = 53\%$; $P_Q = 0.02$). Visual inspection of the funnel plot (Supplementary Figure 66) and Egger's regression test showed no significant small study effects ($P_{Egger} = 0.18$). No subgroup analysis showed significant between-group differences (study design [p = 0.60; Supplementary Figure 67], risk of bias [p = 0.67; Supplementary Figure 68], corticosteroid regimen [p = 0.16; Supplementary Figure 69], corticosteroid dosage [p = 0.69; Supplementary Figure 70] and time to administration [p = 0.69; Supplementary Figure 71]), and meta-regression analyses showed no correlation between the treatment effect and proportion of severe patients (p = 0.42; Supplementary Figure 72) and baseline mean SOFA score (p = 0.95; Supplementary Figure 73).

Nosocomial infection

Twenty-seven studies [30,85,89,90,96,104,105,107,111,113,114,116,120,129,130,134,140,142,144,149,150,153,155,156,158,168,170] with 9483 patients reported incidence of nosocomial infection (Figure 8D) with a pooled OR of 1.06 (95% CI: 0.82– 1.37). There was significant and moderate heterogeneity ($I^2 = 47\%$; $P_Q < 0.01$). Visual inspection of the funnel plot (Supplementary Figure 74) and Egger's regression test showed no significant small study effects ($P_{Egger} = 0.99$). No subgroup analysis showed significant between-group differences (study design [p = 0.89; Supplementary Figure 75], risk of bias [p = 0.27; Supplementary Figure 76], corticosteroid regimen [p = 0.70; Supplementary Figure 77], corticosteroid dosage [p = 0.25; Supplementary Figure 78] and time to administration [p = 0.06; Supplementary Figure 79]), and meta-regression analyses showed no correlation between the treatment effect and proportion of severe patients (p = 0.72; Supplementary Figure 80) and baseline mean SOFA score (p = 0.24; Supplementary Figure 81).

Quality of evidence

The summary of findings for primary outcomes is tabulated in Table 1.

Discussions

Main findings

The current systematic review and meta-analysis included 10 RCTs and 85 observational studies involving 42,205 hospitalized COVID-19 patients, making it the most comprehensive systematic review available on the use of corticosteroid for treating COVID-19. We found that the use of corticosteroid regimens was not significantly associated with changes in length of ICU hospitalization, nor incidence of mechanical ventilation.

While the overall meta-analysis showed that corticosteroids had no impact on length of stay, the pooled MD from RCT studies showed that corticosteroids may significantly increase the length of stay by 3.62 days, while the 95% CI for the pooled MD from observational studies narrowly crossed the line of no effect. The overall meta-analysis also showed that corticosteroids had no impact on mortality nor incidence of ICU admission, but subgroup analysis by corticosteroid type showed that methylprednisolone may confer a significant survival benefit with a 38% reduction in the odds of death compared with standard of care, while other regimens such as dexamethasone and hydrocortisone conferred no significant survival benefit. Additionally, methylprednisolone may be associated with a significant 47% reduction in the incidence of ICU admission. Across mechanically ventilated patients, we found that corticosteroids significantly reduced length of mechanical ventilation. However, corticosteroids may

Table 1. Summary of findings, corticosteroid regimens compared with standard of care/adjuvant therapies for the management of hospitalized COVID-19 patients.

Outcomes	Relative effect (95%	Anticir	ated absolute effects (9	5% CI)†	Patients n (studies n)	Quality of evidence
outcomes	CI)	Risk without corticosteroids	Risk with corticosteroids	Risk difference (95% Cl)	radents, in (statics, ii)	(GRADE)
Time to negative test conversion	-	The mean time in the control group was 15 days	-	MD 1.80 more days (0.48 more to 3.11 more)	2,785 (1 RCT, 18 OSs)	$\oplus \bigcirc \bigcirc \bigcirc$ Very low ^{‡, §, ¶}
Length of stay	-	The mean time in the control group was 16 days	-	MD 1.60 more days (0.20 fewer to 3.40 more)	10,454 (3 RCTs, 27 OSs)	⊕⊜⊜⊜ Very low ^{‡,§,} ¶,#
ICU length of stay	-	The mean time in the control group was 12 days	-	MD 0.01 fewer days (3.50 fewer to 3.47 more)	3,233 (1 RCT, 6 OSs)	⊕⊖⊖⊖ Very low ^{‡,} §,¶,#
Incidence of ICU admission	OR 0.87 (0.52–1.47)	121 per 1000	107 per 1000 (67–168)	14 fewer per 1000 (54 fewer to 47 more)	10,391 (2 RCTs, 15 OSs)	⊕⊖⊖⊖ Very low ^{‡,§,} ¶,#
Mortality	OR 1.00 (0.84–1.20)	235 per 1000	235 per 1000 (205–269)	0 fewer per 1000 (30 fewer to 34 more)	40,623 (10 RCTs, 76 OSs)	⊕⊖⊖⊖ Very low ^{‡,§,} ¶,#
Incidence of mechanical ventilation	OR 1.19 (0.81–1.76)	139 per 1000	161 per 1000 (116–221)	22 more per 1000 (23 fewer to 82 more)	20,453 (4 RCTs, 33 OSs)	⊕⊜⊜⊜ Very low ^{‡,} §,¶,#
Length of mechanical ventilation	-	The mean time in the control group was 6 days	-	MD 1.71 fewer days (2.54 fewer to 0.88 fewer)	977 (1 RCT, 6 OSs)	$\oplus \bigcirc \bigcirc \bigcirc$ Very low ^{‡, §}
Incidence of adverse events	OR 1.89 (0.98–3.63)	362 per 1000	517 per 1000 (512–828)	155 more per 1000 (5 fewer to 311 more)	3,409 (1 RCT, 8 OSs)	⊕⊖⊖⊖ Very low ^{‡,§,#}
Incidence of severe adverse events	OR 0.87 (0.47–1.62)	57 per 1000	50 per 1000 (27–89)	7 fewer per 1000 (30 fewer to 32 more)	1,142 (4 RCTs, 2 OSs)	⊕⊕⊕⊖ Moderate [#]
Incidence of hyperglycemia	OR 2.45 (1.08–5.56)	99 per 1000	212 per 1000 (106–379)	113 more per 1000 (7 more to 280 more)	2,898 (3 RCTs, 8 OSs)	⊕⊕⊖⊖ Low ^{‡,} §,††
Incidence of nosocomial infections	OR 1.06 (0.82–1.37)	150 per 1000	158 per 1000 (126–195)	8 more per 1000 (24 fewer to 45 more)	9,483 (4 RCTs, 23 OSs)	⊕⊖⊖⊖ Very low ^{‡,§,#}

Bold values represent the most important metrics.

GRADE Working Group quality of evidence rating [70].

High quality: We are very confident that the true effect lies close to that of the estimate of the effect.

Moderate quality: We are moderately confident in the effect estimate; the true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different.

Low quality: Our confidence in the effect estimate is limited; the true effect may be substantially different from the estimate of the effect.

Very low quality: We have very little confidence in the effect estimate; the true effect is likely to be substantially different from the estimate of effect.

[†]The risk in the intervention group (and its 95% CI) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).

[±]Quality of study was rated as low prior to downgrading or upgrading as a majority of the included studies were observational studies.

§Downgraded due to study limitations; a majority of included studies were rated as having serious or critical risk of bias according to ROBINS-I and/or RoB2.

[¶]Downgraded due to inconsistency; significant and severe heterogeneity was observed in the analysis.

[#]Downgraded due to imprecision; confidence intervals could not rule out the possibility of no effect (crosses null).

^{††}Upgraded due to a large magnitude of effect.

GRADE: Grading of recommendations, assessment, development and evaluation; MD: Mean difference; OR: Odds ratio; OS: Observational study; RCT: Randomized controlled trial.

also significantly increase time to negative conversion of SARS-CoV-2 tests, which could suggest that corticosteroid administration hinders viral clearance.

In terms of safety, the use of corticosteroids was not associated with a significant increase in the odds of adverse events or severe adverse events. However, the pooled OR from observational studies showed that corticosteroids may increase the odds of adverse events by 117%. We also found a significant difference in the odds of serious adverse events between studies using high fixed-dose methylprednisolone versus low fixed-dose hydrocortisone, with methylprednisolone being associated with a significant, 17% reduction in the odds of serious adverse events while hydrocortisone was associated with a significant 202% increase in the odds of serious adverse events. While assessing the odds of hyperglycemia and nosocomial infection – both common adverse events associated with corticosteroids – we found that corticosteroids significantly increased the odds of hyperglycemia by 145%, however, there were no significant increases in nosocomial infection associated with corticosteroids.

Clinical implications for corticosteroid efficacy

In this review, we found that corticosteroids had no protective effect against the incidence and duration of ICU hospitalization nor incidence of progression to mechanical ventilation. These findings are consistent with previous studies assessing the use of corticosteroids in other coronavirus infections such as SARS-CoV and MERS-CoV [178], demonstrating that corticosteroids may not be effective for preventing progression from mild-and-moderate COVID-19 to severe disease, nor is it effective for managing ICU-hospitalized patients. In fact, corticosteroids may significantly increase the length of stay for hospitalized patients (according to pooling of RCT studies) and increase time to viral clearance, potentially placing more strain on limited hospital resources and increasing the risk of viral transmission. As previously speculated, using immunosuppressive therapies indiscriminately for viral infections such as COVID-19 may complicate disease courses and perpetuate illness [179].

Interestingly, while we found that corticosteroids had no significant impact on the incidence of mortality and ICU admission, methylprednisolone appears to have an exceptional efficacy for reducing the odds of death compared with other corticosteroids such as dexamethasone and hydrocortisone. Previous studies, including a triple-blinded RCT [180] and an observational cohort study [125], also found that methylprednisolone may be more efficacious at reducing mortality compared with dexamethasone. In addition, when the meta-analysis is restricted to studies using methylprednisolone only, it was found that methylprednisolone may also significantly reduce incidence of ICU admissions. It was hypothesized that these observations were due to methylprednisolone having greater lung tissue penetration compared with other corticosteroids [181,182], making methylprednisolone a more optimal corticosteroid for treating hyperinflammatory reactions to SARS-CoV-2 and respiratory complications. However, we suspect that over-precision associated with the large quantity of observational studies in the methylprednisolone subgroup may have contributed to our finding as well [183]. Thus we cannot definitively conclude that methylprednisolone is superior compared with other corticosteroids for reducing mortality or ICU admission. Currently, most published RCTs focus on dexamethasone and hydrocortisone [28,29], but less frequently on methylprednisolone. Further investigations using large scale, high-quality RCTs are required to establish the potential impact of methylprednisolone on mortality and ICU admission.

A main finding of our review is that corticosteroids significantly reduced the length of mechanical ventilation in critically ill patients experiencing respiratory failure, while it was not effective in decreasing the odds of mechanical ventilation incidence. Mechanical ventilation is often required in COVID patients experiencing ARDS, defined as a reduction in the area of normoventilated lung due to inflammatory pulmonary edema [184]. While pulmonary edema is one of the main drivers of COVID-related mortality [185,186], prolonged mechanical ventilation is also associated with a higher risk of developing complications such as ventilator-associated pneumonia [187], death, increased hospital and ICU lengths of stay, and increased cost [187,188]. Thus, the use of corticosteroids may substantially improve the outcome of critically ill COVID-19 patients who require mechanical ventilation, as well as conserving valuable ventilator assets during COVID surges.

Compared with previous meta-analyses, the current review did not identify a robust association between the use of corticosteroids and mortality, unlike smaller systematic reviews conducted earlier during the pandemic [58,59,62]. However, our findings regarding prolonged duration of viral clearance and reduced length of mechanical ventilation is strongly supported by the findings of previous meta-analyses [60,61]. In summary, our results suggest that the role of corticosteroids for treating COVID-19 may be limited to severe and critically ill patients receiving mechanical ventilation. It may not be effective for reducing mortality or improving other efficacy outcomes for other patient groups. The effect of methylprednisolone on the odds of mortality requires further investigation and clarification. Last, while corticosteroids may be effective in certain patient groups, it is currently unknown whether it is related to SARS-CoV-2, its treatment, or both. Further research may be warranted.

Clinical implications for corticosteroid safety

Similar to previous meta-analyses [58,62], we did not identify any significant associations between the use of corticosteroids and the odds of all-cause adverse events and all-cause severe adverse events. However, excluding the CoDEX trial [30] from the meta-analysis resulted in a significantly increased odds of adverse events based on the treatment effects of the remaining observational studies. Given that the patient characteristics are similar between the CoDEX trial and patients in the observational study, especially as both the CoDEX trial and the observational studies largely excluded mild and moderate patients, a possible explanation of this finding is due to the use of corticosteroids in a subset of control patients in the CoDEX trial, as well as the open-label design which may lead to biased reporting of adverse events in the experimental group. The CoDEX investigators stated that these factors

may have shifted the results toward null [30]; consequently, the inclusion of the CoDEX trial would have shifted our pooled results toward null as well.

Additionally, we found that the pooled results from studies using high fixed-dose methylprednisolone yielded a significantly decreased odds of severe adverse events, while studies using low fixed-dose hydrocortisone yielded a significantly increased odds of severe adverse events. As there was no heterogeneity within each of the aforementioned subgroups, and the studies included in the subgroups (including 3 RCTs and 1 propensity-matched cohort study) were unlikely to be substantially biased; we speculate that the differences in corticosteroid doses may explain this observation. It is possible that the high-dose methylprednisolone therapy was more efficacious at controlling COVID symptoms and undesirable effects from adjuvant medications than low-dose hydrocortisone, thus resulting in less reported adverse events. Overall, a majority of the adverse events reported by the subgroup studies were deemed to be unrelated to the intervention [29,87,89,104]. Nevertheless, the sample size supporting this observation is low, and further investigations are needed to confirm this finding.

In this review, we assessed the odds of hyperglycemia and nosocomial infections, two of the most common types of adverse events in patients undergoing systemic corticosteroid treatments [189]. Our review found that the use of corticosteroids is associated with a significant increase in the odds of hyperglycemia compared with control, indicating a need for glycemic monitoring for patients receiving corticosteroid therapy [190]. Interestingly, we did not identify a significant correlation between corticosteroid use and the odds of nosocomial infection, contrary to expectations due to corticosteroids' immunosuppressive properties. This is most likely due to the use of concomitant antimicrobials in COVID-19 patients to prevent superinfections. As the reporting of antibiotic and antifungal use among our included studies were extremely inconsistent, we could not properly evaluate the impact of corticosteroids on the odds of nosocomial infections.

Strengths & limitations

This systematic review and meta-analysis has several strengths. First, we conducted several subgroup analyses to assess sources of heterogeneity and to study the impact of study methodologies on the treatment effect, including comparing the pooled effect from randomized to non-randomized studies and studies with low-moderate risk of bias versus studies with high risk of bias. Secondly, we followed the most up-to-date systematic review guidelines and methods, such as PRISMA 2020, RoB2, and GRADE. Moreover, the large sample size in our analyses improved the power and precision of our study compared with previous systematic reviews.

The major limitation of this study is the inclusion of observational studies. Similar to our previous systematic reviews [25,26], a majority of our included observational studies were rated as having a critical or serious risk of bias according to ROBINS-I, highlighting the poor publication quality associated with pandemic-related literature [191]. The inclusion of observational studies in meta-analyses was also associated with large treatment effects [192,193] and overly precise estimates [194,195], potentially introducing biases into our significant findings. However, it must be noted that our subgroup analysis by study design showed that the findings from observational studies were largely in agreement with findings from RCTs for a majority of the outcomes. In addition, while we conducted numerous subgroup analyses and meta-regressions to explore the sources of heterogeneity and possible factors that may affect the efficacy of corticosteroids, we could not provide insight on the effect of virus variants, the general care provided, vaccination status and underlying conditions on the efficacy of corticosteroids over the course of the pandemic due to a paucity of data. With rapidly changing COVID-19 treatment strategies and the spread of the Omicron variant, future studies may wish to consider the potential impact of these factors.

Conclusion

This comprehensive systematic review and meta-analysis showed that the use of corticosteroids was not significantly associated with changes in length of ICU hospitalization, incidence of ICU admission, incidence of mortality, nor incidence of mechanical ventilation. Additionally, corticosteroids may increase length of stay (based on RCT data only) and increase the duration of viral clearance. However, corticosteroids were significantly associated with reduction in the length of mechanical ventilation; thus, it may confer beneficial effects for critically ill patients requiring mechanical ventilation. There was also evidence suggesting that methylprednisolone may be efficacious in reducing mortality and incidence of ICU admission; however, these findings should be further assessed using RCTs.

In terms of safety, corticosteroids were not associated with increased odds of adverse events or severe adverse events, although these findings need to be interpreted with caution due to potential confounding factors such as corticosteroid dosage. Corticosteroids were also not associated with increased odds of nosocomial infections, albeit the use of adjuvant antimicrobials complicates the interpretation of this finding. Last, corticosteroids use was associated with significantly increased odds of hyperglycemia, indicating a need for glycemic monitoring in COVID-19 patients receiving corticosteroid treatments.

Summary points

- Corticosteroids are a category of drugs repurposed for the management of COVID-19 patients with conflicting evidence regarding its efficacy.
- In this systematic review and meta-analysis, data from 10 randomized controlled trials (RCTs) and 85 observational studies (n = 42,205) were analyzed using random-effects meta-analyses to assess its efficacy and safety for treating hospitalized COVID-19 patients.
- A majority of included RCTs were rated as having some concerns in terms of risk of bias according to RoB2, while a majority of included observational studies were rated as having serious or critical risk of bias according to ROBINS-I.
- The use of corticosteroids was not significantly associated with changes in length of ICU hospitalization, incidence of ICU admission, nor incidence of mechanical ventilation.
- The use of corticosteroids was significantly associated with increased length of stay (based on RCT data only) and increased time to negative SARS-CoV-2 test; additionally, the use of corticosteroids was associated with reduced length of mechanical ventilation.
- In terms of safety, the use of corticosteroids was not associated with increased odds of adverse events, serious adverse events, or nosocomial infections; however, the use of corticosteroids was associated with increased odds of hyperglycemia.
- Some evidence suggests that methylprednisolone may reduce mortality and incidence of ICU admission, and high-dose methylprednisolone pulses may result in reduced incidence of serious adverse events compared with other regimens such as low-dose hydrocortisone; however more research is needed to confirm these findings.
- Current evidence supports the use of corticosteroids in the management of critically ill COVID-19 patients requiring corticosteroids, and glycemic monitoring is required for patients undergoing treatments with corticosteroids.

Supplementary data

To view the supplementary data that accompany this paper please visit the journal website at: www.futuremedicine.com/doi/ suppl/10.2217/fvl-2021-0244

Author contributions

F Zhou was responsible for conception and design of the work, as well as for supervising and contributing to all parts of the data acquisition process (including article screening, full-text retrieval, data extraction and risk of bias analyses), drafting of the final manuscript and revised the final manuscript critically for important intellectual content. J Deng was responsible for conception and design of the work, as well as for search strategy development, conducting database searches, performing data analyses (including meta-analyses, meta-regressions, subgroup analyses and GRADE ratings) and drafting of the final manuscript. K Heybati, QK Zuo, S Ali, W Hou, CY Wong, HB Ramaraju, O Chang, T Dhivagaran and Z Silver contributed to all parts of the data acquisition process (including article screening, full-text retrieval, data extraction and risk of bias analyses) and revised the final manuscript critically for important intellectual content. All authors have given final approval for the final version of this manuscript to be submitted for publication and all authors agree to be held accountable for all aspects of the work in ensuring that questions related to the accuracy or integrity of any part of the work are appropriately investigated and resolved.

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Ethical conduct of research

There are no relevant ethical disclosures as only published, aggregate patient data was used for this study. Aggregate patient data extracted by study authors are available upon reasonable request.

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